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Predicting the impact of non-pharmaceutical interventions against COVID-19 on *Mycoplasma pneumoniae* in the United States

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A R T I C L E I N F O

Dataset link: [https://github.com/parksw3/myc](https://github.com/parksw3/mycoplasma_pred) [oplasma_pred](https://github.com/parksw3/mycoplasma_pred)

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A B S T R A C T

The introduction of non-pharmaceutical interventions (NPIs) against COVID-19 disrupted circulation of many respiratory pathogens and eventually caused large, delayed outbreaks, owing to the build up of the susceptible pool during the intervention period. In contrast to other common respiratory pathogens that re-emerged soon after the NPIs were lifted, longer delays (*>* 3 years) in the outbreaks of *Mycoplasma pneumoniae* (Mp), a bacterium commonly responsible for respiratory infections and pneumonia, have been reported in Europe and Asia. As Mp cases are continuing to increase in the US, predicting the size of an imminent outbreak is timely for public health agencies and decision makers. Here, we use simple mathematical models to provide robust predictions about a large Mp outbreak ongoing in the US. Our model further illustrates that NPIs and waning immunity are important factors in driving long delays in epidemic resurgence.

1. Introduction

Mycoplasma pneumoniae (Mp) is the most commonly detected bacterium for lower tract respiratory infections in children and adults (Waites and [Talkington,](#page-5-0) [2004](#page-5-0); [Jain](#page-5-1) et al., [2015a,](#page-5-1)[b](#page-5-2); [Bajantri](#page-5-3) et al., [2018\)](#page-5-3). Mp pneumonia can affect a patient for a long period due to its long incubation period and long prodromal duration of symptoms, especially among children and high risk individuals. For instance, outbreaks in schools have resulted in increased hospitalization, ventilator-associated pneumonia, severe extrapulmonary disease (e.g. Stevens-Johnson syndrome) and have been shown to last for several months ([Walter](#page-5-4) et al., [2008;](#page-5-4) [Olson](#page-5-5) et al., [2015\)](#page-5-5). Increasing levels of macrolide-resistant Mp strains further highlight the importance of appropriate diagnosis and treatment ([Pereyre](#page-5-6) et al., [2016\)](#page-5-6).

Multiannual cycles in Mp outbreak patterns have been observed in many countries, with large outbreaks occurring every 3–7 years [\(Kim](#page-5-7) et [al.](#page-5-7), [2009;](#page-5-7) [Brown](#page-5-8) et al., [2016](#page-5-8)). Several potential mechanisms have been proposed to explain the observed epidemic cycles, including waning immunity ([Omori](#page-5-9) et al., [2015](#page-5-9)) and strain dynamics [\(Kenri](#page-5-10) et al., [2008;](#page-5-10) [Zhang](#page-5-11) et al., [2019\)](#page-5-11). More parsimoniously, multiennial epidemic dynamics can also arise from simple, seasonally-forced immunizing epidemic dynamics, especially given low transmission rates ([Earn](#page-5-12) et al., [2000;](#page-5-12) [Keeling](#page-5-13) et al., [2001](#page-5-13)).

Along with other endemic viruses and bacteria, the circulation of Mp was disrupted in 2020 due to non-pharmaceutical interventions (NPIs) that were introduced to prevent the transmission of SARS-CoV-2 ([Boyanton](#page-5-14) Jr. et al., [2024](#page-5-14)). The disruption of the expected epidemiological curve adds challenges to predicting future Mp outbreaks. In contrast to other common respiratory pathogens that re-emerged within a year after NPIs were lifted, the re-emergence of Mp outbreaks has been delayed for more than 3 years in Europe and Asia [\(Sauteur](#page-5-15) et al., [2024\)](#page-5-15). A long delay in epidemic resurgence is particularly alarming because it can allow for a build up of a large susceptible population, increasing the risk of a large outbreak as the infection resurges ([Baker](#page-5-16) et [al.,](#page-5-16) [2020](#page-5-16)).

As Mp infections are beginning to increase rapidly in the US ([Fig.](#page-1-0) [1A](#page-1-0)), predicting the size of an impending outbreak is critical for public health agencies and decision makers. Here, we present a modeling analysis of Mp outbreaks in the US using data from 2015

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Fig. 1. Summary of model fits to Mycoplasma pneumoniae incidence proxy in the US,2015-2024. (A) Comparisons of observed (points) and fitted (line) trajectories of incidence proxy for Mp infections. (B) Estimated non-periodic, time-varying transmission term, representing relative transmission δ following the introduction of NPIs; for example, 0.5 corresponds to a 50% reduction in transmission. (C) Estimated periodic transmission term $\beta_{\rm{seas}}(t)$, representing seasonal transmission rate. Red lines and shaded regions represent the estimated posterior median and corresponding 95% and 50% credible intervals. (For interpretation of the references to color in this figure legend, the reader is referred to the web version of this article.)

onward and predictions for potential upcoming outbreaks and subsequent epidemic dynamics. We also explore counterfactual scenarios to tease apart factors that contributed to long delays in the resurgence of Mp outbreaks.

2. Methods

2.1. Data

We analyzed over 2.57 million BIOFIRE® Respiratory (RP1.7, RP2, and RP2.1) Panel (bioMérieux, Salt Lake City, Utah) multiplex PCR test results, 8678 of which had a positive detection of Mp [\(Poritz](#page-5-17) et al., [2011;](#page-5-17) [Leber](#page-5-18) et al., [2018](#page-5-18); [Creager](#page-5-19) et al., [2020](#page-5-19)). These deidentified patient test results were captured from 127 US inpatient and outpatient facilities from January 1, 2015 to June 29, 2024 using BIOFIRE® Syndromic Trends, a cloud-based pathogen surveillance network for BIOFIRE® instruments. Further details regarding the dataset and test result collection and interpretation methods have previously been published [\(Meyers](#page-5-20) et al., [2018](#page-5-20)). We also analyzed time-series data of influenza-like illnesses (ILIs), which were downloaded the FluView website (<https://gis.cdc.gov/grasp/fluview/fluportaldashboard.html>), and time-series data of Google Mobility for the US, which were taken from a previous publication [\(Park](#page-5-21) et al., [2024\)](#page-5-21).

2.2. Incidence proxy

As discussed previously, test positivity can give a biased view of disease circulation patterns ([Goldstein](#page-5-22) et al., [2011;](#page-5-22) [Kissler](#page-5-23) et al., [2020](#page-5-23);

[Park](#page-5-21) et al., [2024](#page-5-21)). Specifically, since test positivity in hospitals reflects prevalence of infection among individuals exhibiting respiratory symptoms, an increase (or decrease) in circulations of other respiratory diseases can cause test positivity of Mp infections to decrease (or increase). To mitigate this bias, we calculated an incidence proxy for Mp infections by multiplying weekly percentages of MP positives with weekly percentages of weighted ILIs. Limitations and assumptions to using this incidence proxy are discussed in [Goldstein](#page-5-22) et al. ([2011](#page-5-22)). Comparisons of three time series (Mp positivity, ILI positivity, and incidence proxy) are provided in Supplementary Figure S1.

2.3. Mathematical modeling of Mp outbreaks

To infer epidemic dynamics of Mp infections in the US, we fitted a seasonally-forced, deterministic Susceptible–Infectious–Recovered– Susceptible (SIRS) model to Mp time series data [\(Dushoff](#page-5-24) et al., [2004](#page-5-24); [Shaman](#page-5-25) et al., [2010](#page-5-25)). The SIRS model consists of three compartments, each representing the current infection status of an individual: Susceptible, Infected, and Recovered. We assumed a homogeneously mixed population, where infected individuals (I) transmit infection to susceptible individuals (S) at rate $\beta(t)$ and recover from infection at rate $\gamma = 1/3$ weeks [\(Omori](#page-5-9) et al., [2015\)](#page-5-9). Recovered individuals (R) were assumed to return to a susceptible state at rate v , which we estimated by fitting the model to data. Birth and death rates $\mu = 1/80$ years were assumed to be equal to fix the population size $(S + I + R = 10⁶)$. The model was discretized using the Euler scheme outlined in [He](#page-5-26) et al. ([2010\)](#page-5-26) at a weekly time scale $\Delta t = 1$ week:

$$
\Delta S(t) = \left[1 - \exp(-(\beta(t)I(t) + \mu)\Delta t)\right]S(t - \Delta t) \tag{1}
$$

S.W. Park et al.

$$
N_{SI}(t) = \frac{\beta(t)I(t)\Delta S(t)}{\beta(t)I(t) + \mu}
$$
\n(2)

$$
\Delta I(t) = \left[1 - \exp(-(\gamma + \mu)\Delta t)\right]I(t - \Delta t)
$$
\n(3)

$$
N_{IR}(t) = \frac{\gamma \Delta I(t)}{\gamma + \mu} \tag{4}
$$

$$
\Delta R(t) = \left[1 - \exp(-(v + \mu)\Delta t)\right] R(t - \Delta t)
$$
\n(5)

$$
N_{RS}(t) = \frac{\mu_{SN}(t)}{v + \mu} \tag{6}
$$

$$
S(t) = S(t - \Delta t) + \mu - \Delta S(t) + N_{RS}(t)
$$
\n(7)

$$
I(t) = I(t - \Delta t) - \Delta I(t) + N_{SI}(t)
$$
\n(8)

$$
R(t) = R(t - \Delta t) - \Delta R(t) + N_{IR}(t)
$$
\n(9)

where $\Delta X(t)$ represents the total number of individuals leaving the compartment X at time *t*, and $N_{XY}(t)$ represents the number of individuals leaving the compartment X to enter the compartment Y at time t . We further calculated the incidence $C(t)$ by keeping track of the number of new infections in each week:

$$
C(t) = N_{SI}(t),\tag{10}
$$

which was fitted to estimated incidence proxy using a log-normal likelihood:

$$
log(incidence proxy) \sim Normal(log(\rho C(t)), \sigma_{obs}),
$$
\n(11)

where, ρ represents the scaling factor and $\sigma_{\rm obs}$ represents the residual standard deviation. Typically, either Poisson or negative binomial likelihoods are used to model discrete case counts; since incidence proxy is a continuous variable, we relied on the log-normal likelihood instead. For both ρ and $\sigma_{\rm obs}$, we used weakly informative half-normal priors:

$$
\rho \sim \text{Half-Normal}(0, 2) \tag{12}
$$

$$
\sigma_{\rm obs} \sim \text{Half-Normal}(0, 0.5). \tag{13}
$$

In order to capture epidemic dynamics before and after NPIs were introduced, we extended the SIRS model by decomposing the transmission rate $\beta(t)$ into two separate terms: (1) a periodic term with a period of 1 year $\beta_{\text{seas}}(t)$ that captures seasonal transmission, reflecting epidemic peaks in summer and fall ([Fig.](#page-1-0) [1A](#page-1-0)) and (2) a non-periodic, time-varying term that captures changes in contact patterns after the introduction of COVID-19 NPIs in March, 2020 $\delta(t)$:

$$
\beta(t) = \beta_{\text{seas}}(t)\delta(t),\tag{14}
$$

where δ < 1 corresponds to reduction in transmission due to NPI effects. The periodic term $\beta_{\text{seas}}(t)$ was modeled by estimating a weekly transmission rate for each of 52 weeks, which was modeled using cyclical, random-walk priors to allow for smoothing:

$$
\beta_{\text{seas}}(t) \sim \text{Normal}(\beta_{\text{seas}}(t-1), \sigma) \qquad t = 2 \dots 52 \tag{15}
$$

$$
\beta_{\text{seas}}(1) \sim \text{Normal}(\beta_{\text{seas}}(52), \sigma) \tag{16}
$$

To prevent estimating transmission rates that are unrealistically low or high, we also added a weakly informative prior:

$$
\beta_{\text{seas}}(t) \sim \text{Normal}(1, 0.5). \tag{17}
$$

A half-normal prior was used for the standard deviation of transmission smoothing σ :

$$
\sigma \sim \text{Half-Normal}(0, 0.2). \tag{18}
$$

The non-periodic term was modeled by estimating a separate δ for every four weeks with a normal prior centered around 1:

$$
\delta(t) \sim \text{Normal}(1, 0.2). \tag{19}
$$

We also used a weakly informative prior for the duration of immunity $\tau = 1/v$ to prevent extremely long or short duration of immunity:

 $\tau \sim \text{Normal}(300, 100)$. (20)

Finally, we assumed a uniform prior for the initial proportion susceptible $S(0)$ and a half-normal prior for the initial proportion infected $I(0)$:

$$
S(0) \sim \text{Unif}(0, 1) \tag{21}
$$

$$
I(0) \sim \text{Half-Normal}(0, 0.01) \tag{22}
$$

The model was fitted using rstan ([Carpenter](#page-5-27) et al., [2017;](#page-5-27) [Stan](#page-5-28) De[velopment](#page-5-28) Team, [2024](#page-5-28)) with 4 MCMC chains, each with 4000 iterations. Convergence was assessed based on the lack of warning signs from rstan, indicating: no divergent chains; no iterations exceeding maximum tree depth; sufficiently high Bayesian Fraction of Missing Information; sufficiently high effective samples sizes (*>* 400 for every parameter); and sufficiently low Rhat (*<* 1*.*01). The resulting posterior distribution was used to predict future Mp outbreaks by projecting the model forward.

As a sensitivity analysis, we also tried fitting the SIR model, which assumes that infection provides life-long immunity. The SIR model was implemented by setting the immune waning rate to zero, $v = 0$. This model was also fitted using the same priors and likelihood function with 4 MCMC chains, each with 6000 iterations.

2.4. Evaluating the impact of NPIs and waning immunity on the delays in re-emergence of Mp outbreak

It is currently unknown what factors contributed to long delays in re-emergence of the Mp outbreak. To evaluate the potential impact of NPIs and waning immunity on the timing of Mp re-emergence, we explored counterfactual scenarios by performing sensitivity analyses on how different assumptions about NPI effects and waning immunity can affect the timing of Mp outbreak resurgence; for simplicity, we tested the sensitivity of epidemic dynamics to each assumption independently (i.e., by varying one variable at a time) and did not explore their joint effects. First, the strength of NPIs were modified by taking the estimated fold changes in transmission $\delta - 1$ and scaling it by a factor of θ such that the resulting transmission rate corresponds to

$$
\beta(t) = \beta_{\text{seas}}(t)[1 + \theta(\delta(t) - 1)].\tag{23}
$$

Second, the duration of NPIs were modified by assuming $\delta = 1$ after the end date t_{end} :

$$
\beta(t) = \begin{cases} \beta_{\text{seas}}(t)\delta(t) & t < t_{\text{end}}\\ \beta_{\text{seas}}(t) & t \ge t_{\text{end}}. \end{cases} \tag{24}
$$

We varied t_{end} between 2020–2024. We also tested how the timing of NPI introduction, relative to the timing of multiannual cycles of Mp outbreak, by shifting $\delta(t)$ by 1–3 years to allow earlier introduction of NPIs. Finally, we varied the duration of immunity $1/v$ between 5– 25 years to assess how waning immunity contributes to the timing of Mp outbreak resurgence. For illustration purposes, all other parameters are set to median values from the posterior throughout the analysis.

3. Results

Mathematical modeling of past Mp epidemics. The model reproduced the observed epidemic dynamics for Mp infections, including the seasonal and longer-term (\approx 5 year) epidemic cycles before NPIs were introduced and the subsequent delayed resurgence of the epidemic ([Fig.](#page-1-0) [1](#page-1-0)A). To capture the response of Mp to the pandemic, the model required a strong reduction in transmission $(\delta(t) < 1)$ in 2020 ([Fig.](#page-1-0) [1B](#page-1-0)). The model further estimated sustained reduction in transmission for 2021–2023 [\(Fig.](#page-1-0) [1](#page-1-0)C). While we did not find a statistical correlation between the estimated changes in transmission and mean changes in Google mobility measures $(r = 0.14; 95\% \text{ CI: } -0.03, 0.30)$, the strength of correlation increased with lags, with strongest correlation occurring at a 4-week lag ($r = 0.36$; 95% CI: 0.21-0.50; Supplementary Figure S2). The model demonstrated sinusoidal patterns in seasonal transmission which is typical for respiratory infections—peaking around week 43

Fig. 2. Predictions of future Mycoplasma pneumoniae outbreaks. (A) Predicted changes in the proportion of susceptible individuals. (B) Predicted changes in weekly incidence proxy for Mp infections. Red lines and shaded regions represent the estimated posterior median and corresponding 95% and 50% credible intervals. Points represent the observed incidence proxy. (For interpretation of the references to color in this figure legend, the reader is referred to the web version of this article.)

([Fig.](#page-1-0) [1C](#page-1-0)). The model also estimated the mean duration of immunity to be 15.1 years (95% CI: 11.9 years–18.7 years). Comparisons between posterior and prior distributions are presented in Supplementary Figure S3. The SIR model gave nearly identical estimates for the NPI effects and transmission rates but underestimated the peak of the 2015 outbreak, providing support for the SIRS model (Supplementary Figure S4).

Prediction of future Mp outbreaks. In line with other acute, partially immunizing respiratory pathogens ([Baker](#page-5-16) et al., [2020\)](#page-5-16), the model predicted a large build-up of the susceptible pool during the NPI period [\(Fig.](#page-3-0) [2](#page-3-0)A). This increase in susceptible individuals could lead to a large Mp outbreak, which appears to have begun since the beginning of 2024 ([Fig.](#page-1-0) [1](#page-1-0)A); the model forecasted that the epidemic would peak around the second half of 2024 and start to decline towards the end of 2024 [\(Fig.](#page-3-0) [2](#page-3-0)B), reflecting a decrease in transmission rate in winter ([Fig.](#page-1-0) [1](#page-1-0)C). The peak incidence during this outbreak is estimated to be 4.9 (95% CI: 4.0–6.0) times larger than previous peaks. The model further predicted that the outbreak beginning in 2024 would cause a large depletion of the susceptible pool ([Fig.](#page-3-0) [2](#page-3-0)A), which would lead to another delayed outbreak after several years [\(Fig.](#page-3-0) [2](#page-3-0)B).

Impact of NPI pattern and waning immunity on the timing of Mp outbreak resurgence. The model predicted that a strong reduction in transmission due to NPIs is critical to explaining long delays in the Mp outbreak resurgence. A smaller reduction in transmission [\(Fig.](#page-4-0) [3](#page-4-0)A) would have led to more persistent epidemics as well as earlier resurgence of Mp outbreaks [\(Fig.](#page-4-0) [3](#page-4-0)B). Interestingly, the model still predicted a large outbreak in 2023 even when we assumed that NPI effects were 50%–75% weaker than what we originally estimated, demonstrating the risk of a susceptible build up. Similarly, assuming shorter duration of NPIs ([Fig.](#page-4-0) [3C](#page-4-0)) led to earlier resurgence [\(Fig.](#page-4-0) [3](#page-4-0)D). However, even if transmission rates were to return to pre-pandemic values by the end of 2020, the model predicted that the resurgence would not be observed until the middle of 2022 [\(Fig.](#page-4-0) [3D](#page-4-0)). We also find that the timing of NPIs, relative to the timing of multiannual cycles of Mp outbreak, can affect the timing of resurgence (Supplementary Figure S5). Specifically, when NPIs are introduced after a major outbreak, a large susceptible depletion caused by the major outbreak can drive a longer delay until the resurgence. Instead, if NPIs were introduced 1 or 2 years earlier (thus before the major outbreak that occurred in the winter of 2019), the duration of honeymoon period would have been shorter. Finally, we found that the duration of immunity was also a key

factor in determining the delays in Mp outbreak resurgences ([Fig.](#page-4-0) [3E](#page-4-0)): faster waning of immunity would have caused a faster build up of the susceptible pool, leading to an earlier Mp outbreak.

4. Discussion

In this study, we investigated the impact of NPIs on future Mp outbreaks in the US. By fitting a mathematical model to syndromic surveillance data, we predicted that a large Mp outbreak is imminent with a peak expected before the end of 2024. The upcoming outbreak is expected to be much larger than past outbreaks and may last until the end of 2025; note, however, that there is currently substantial uncertainty associated with longer-term dynamics, including the duration of 2024–2025 outbreak and the timing of the subsequent outbreaks. By the beginning of 2024, we could already observe what seems to be the beginning of this outbreak with an increasing rate of Mp detections throughout May and June 2024. This prediction should alert clinicians and health care systems to be prepared for an increase in cases of pneumonia and potentially more rare presentations of Mp, such as reactive infectious mucocutaneous eruption (RIME) and encephalitis.

Our analysis highlighted the importance of a strong reduction in transmission in explaining the delayed resurgence of Mp outbreaks. Specifically, we estimated $\approx 50\%$ reduction in transmission in 2020, which is considerably larger than the reduction in transmission estimated for other respiratory pathogens. For example, [Baker](#page-5-16) et al. ([2020\)](#page-5-16) estimated a 20% reduction in RSV transmission in the beginning of 2020. Future studies should explore whether the epidemiology and life history of Mp infections cause its transmission to be more susceptible to behavioral changes. Our analyses also suggest that the duration of NPIs, relative timing of NPIs, and the duration of immunity against MP infections also contributed to the long delays in resurgence.

There are several limitations to our study. The dataset inherently contains a subset of the US population that has presented to facilities that utilizes the BIOFIRE® Respiratory Panel and participates in the BIOFIRE® Syndromic Trends network. As with many surveillance network data, this assumes inherent selection bias for less healthy individuals than the general population, therefore introducing the limitation of representativeness of the dataset. Additionally, due to the nature of deidentified data collection within the dataset, we are not able to conduct analyses based on patient or facility characteristics or account for variability in diagnostic testing algorithms between

Fig. 3. Impact of strength and duration of NPIs on the timing of Mp outbreak resurgence. (A, C) Assumed values for the relative transmission term δ following the introduction of NPIs by varying the strength (A) and duration (C) of NPIs. (B, D) Resulting epidemic dynamics across different assumptions about NPIs. (E) Resulting epidemic dynamics across different assumptions about the duration of immunity. The arrows indicate the beginning of the observed outbreak as a reference.

facilities. However, this dataset has evidenced correlation with other surveillance sources, reducing concern regarding representativeness and generalizability ([Meyers](#page-5-20) et al., [2018](#page-5-20)).

We did not account for heterogeneity in Mp dynamics across states as the quantity of data did not allow for reliable model-fitting at a regional scale. We did not include age structure in the model as it would require more data. We also did not account for Mp strain dynamics, which have been hypothesized as another major driver of epidemic cycles [\(Kenri](#page-5-10) et al., [2008;](#page-5-10) [Zhang](#page-5-11) et al., [2019\)](#page-5-11). Finally, our estimate of NPI effects need to be interpreted with care, especially during a period with very little case data; in particular, our estimates of NPI effects will try to capture all other mechanisms that affect infection and transmission, not just those directly caused by NPIs. Despite these limitations, our prediction that a build up of a susceptible pool over the past 4 years will cause a large Mp outbreak is likely qualitatively robust.

So far, there have been limited modeling studies analyzing epidemic dynamics of Mp infections ([Omori](#page-5-9) et al., [2015;](#page-5-9) [Zhang](#page-5-11) et al., [2019](#page-5-11)). Our study reinforces recent work on the importance of understanding susceptible dynamics to predict the impact of perturbations to transmission ([Baker](#page-5-16) et al., [2020](#page-5-16); [Park](#page-5-21) et al., [2024](#page-5-21)). As such, our analysis underlines the importance of serological surveillance data to capture the build up of the susceptible pool to foresee future outbreaks ([Mina](#page-5-29) et [al.](#page-5-29), [2020;](#page-5-29) [Nguyen-Tran](#page-5-30) et al., [2022\)](#page-5-30). Our study also underlines the potential of BIOFIRE® Syndromic Trends network as a powerful surveillance measure ([Park](#page-5-31) et al., [2021](#page-5-31)).

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CRediT authorship contribution statement

Sang Woo Park: Writing – review & editing, Writing – original draft, Software, Methodology, Investigation, Formal analysis, Conceptualization. **Brooklyn Noble:** Writing – review & editing, Investigation, Data curation, Conceptualization. **Emily Howerton:** Writing – review & editing, Investigation, Conceptualization. **Bjarke F. Nielsen:** Writing – review & editing, Investigation, Conceptualization. **Sarah Lentz:** Writing – review & editing, Investigation. **Lilliam Ambroggio:** Writing – review & editing, Investigation, Conceptualization. **Samuel Dominguez:** Writing – review & editing, Investigation, Conceptualization. **Kevin Messacar:** Writing – review & editing, Investigation, Conceptualization. **Bryan T. Grenfell:** Writing – review & editing, Writing – original draft, Investigation, Conceptualization.

Declaration of competing interest

The authors declare the following financial interests/personal relationships which may be considered as potential competing interests: Sang Woo Park reports financial support was provided by Life Sciences Research Foundation. Bjarke F Nielsen reports financial support was provided by Carlsberg Foundation. Kevin Messacar reports financial support was provided by National Institute of Allergy and Infectious Diseases Vaccine Research Center. Brooklyn Noble reports a relationship with bioMérieux that includes: employment. Sarah Lentz reports a relationship with bioMérieux that includes: employment. Lilliam Ambroggio reports a relationship with Pfizer Inc. that includes: funding grants. Samuel Dominguez reports a relationship with BIOFIRE Diagnostics, DelveBio, and Karius that includes: consulting or advisory and funding grants. If there are other authors, they declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

Appendix A. Supplementary data

Supplementary material related to this article can be found online at [https://doi.org/10.1016/j.epidem.2024.100808.](https://doi.org/10.1016/j.epidem.2024.100808)

Data availability

All code is stored in a publicly available GitHub repository ([https:](https://github.com/parksw3/mycoplasma_pred) [//github.com/parksw3/mycoplasma_pred](https://github.com/parksw3/mycoplasma_pred)).

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